

## Lecture 11 — November 21, 2024

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## 1 Outline

**Agenda:**

1. Overview of asymptotic normality (optimality and confidence intervals)
2. Probability recap
3. Delta Method

**Last time:**

1. Prediction confidence intervals in linear model
2. Conformal prediction
3. Risk-controllability
4. Multiple testing
5. FWER (Bonferroni)
6. FDR (Benjamini-Hochberg)

## 2 Asymptotics

**Problem Setting**

Let  $X_i \in \mathcal{X} \sim P$ ,  $i = 1, 2, \dots, n$  be IID data points. We want to estimate  $\theta \in \mathbb{R}^d$  ( $\theta \mapsto P$ ) with an estimator:  $\hat{\theta}_n : \mathcal{X}^n \rightarrow \mathbb{R}^d$  (the form of the estimator is predetermined, but data points are random).

**Goal**

Wish to understand the behavior of  $\hat{\theta}_n$  (limiting behavior):

- A good estimator (theoretical).
- Confidence intervals (CIs) based on  $\hat{\theta}_n$  (fixed on a prediction).

## Convergence

It turns out for most  $\hat{\theta}_n$ ,

$$\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} \mathcal{N}(0, \Sigma), \quad \hat{\theta}_n \sim \mathcal{N}(\theta, \Sigma_n).$$

## Optimality

The smaller  $\Sigma$  is, the better the estimator (in the sense of SPD matrices).

## Inference

Also want to characterize approximate confidence intervals:

$$\hat{\theta}_n^{(j)} \pm 1.96\sqrt{\Sigma_{jj}/n}.$$

## 3 Probability Reminder

**Theorem 1** (CLT). Let  $X_i \in \mathbb{R}^k$ ,  $i = 1, 2, \dots$ , be i.i.d.  $\sim P$ , with finite  $\mu = \mathbb{E}[X_i]$  and  $\Sigma = \mathbb{E}[(X_i - \mu)(X_i - \mu)^\top]$ . Then:

$$\sqrt{n}(\bar{X} - \mu) \xrightarrow{d} \mathcal{N}(0, \Sigma).$$

**Theorem 2** (Continuous Mapping). Let  $Y_n \xrightarrow{d} Y^*$  and  $g$  be a continuous function. Then the followings hold.

1.  $g(Y_n) \xrightarrow{d} g(Y^*)$ , if  $g$  is continuous.
2. If  $Y_n \xrightarrow{p} c$ ,  $g$  is continuous at  $c$ , then  $g(Y_n) \xrightarrow{p} g(c)$ .

**Theorem 3** (Slutsky). Let  $Y_n \xrightarrow{d} Y^*$  and  $Z_n \xrightarrow{p} c$  (where  $c$  is a constant). Then the followings hold.

1.  $Y_n + Z_n \xrightarrow{d} Y^* + c$ .
2.  $Y_n Z_n \xrightarrow{d} Y^* c$ .
3.  $Y_n / Z_n \xrightarrow{d} Y^* / c$ , if  $c \neq 0$ .

**Definition 4** (Uniform Tightness). Let  $Y_n$  be random vectors in  $\mathbb{R}^k$ . The sequence  $\{Y_n\}$  is uniformly tight if  $\forall \epsilon > 0$ ,  $\exists M > 0$  such that:

$$\sup_n \mathbb{P}(\|Y_n\| > M) < \epsilon.$$

Uniform Tightness is an analogy of a bounded deterministic sequence (*sequential compactness*). With such a property, no probability mass is escaping to infinity.

**Definition 5.** Some other convenient definitions/notations are summarized here.

1.  $o_p(1)$  is a sequence  $Y_n \xrightarrow{p} 0$  (where  $Y_n$  is a random vector).
2.  $O_p(1)$  is a sequence that is uniformly tight.
3. For random variables  $R_n$ , we have:
  - $X_n = o_p(R_n)$  if  $X_n = Y_n R_n$ , where  $Y_n = o_p(1)$ .
  - $X_n = O_p(R_n)$  if  $X_n = Y_n R_n$ , where  $Y_n = O_p(1)$ .
4. Analogy in scalar sequences:
  - $o(1)$ : Sequence converges to 0.
  - $O(1)$ : Bounded sequence.

**Proposition 6.** With the definitions above, the following statements hold:

- $o_p(1) + o_p(1) = o_p(1)$ .
- $o_p(1) + O_p(1) = O_p(1)$ .
- $o_p(R_n) = R_n o_p(1)$  by definition.
- $O_p(R_n) = R_n O_p(1)$ .

## 4 Delta Method (Taylor Theorem + Probability)

Delta method is a method to figure out the behavior of functions of sequences. Suppose

$$\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} \mathcal{N}(0, \Sigma),$$

and we want to know the behavior of  $\phi(\hat{\theta}_n)$ , where  $\phi: \mathbb{R}^d \rightarrow \mathbb{R}^m$ .

### Idea: Taylor Expansion

Since  $\hat{\theta}_n \xrightarrow{p} \theta$ ,  $\hat{\theta}_n$  is close to  $\theta$ . By Taylor Expansion, we have

$$\phi(\hat{\theta}_n) \approx \phi(\theta) + \phi'(\theta)(\hat{\theta}_n - \theta),$$

where  $\phi'$  is the Jacobian (gradient transpose). Since this is an affine transform, we have:

$$\phi(\hat{\theta}_n) - \phi(\theta) \xrightarrow{d} \mathcal{N}\left(0, \frac{\phi'(\theta)\Sigma\phi'(\theta)^\top}{n}\right).$$

**Theorem 7** (Delta Method). Let  $\phi: \mathbb{R}^d \rightarrow \mathbb{R}^m$  be differentiable at  $\theta$ . If

$$\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{d} \mathcal{N}(0, \Sigma),$$

then

$$\sqrt{n}\left(\phi(\hat{\theta}_n) - \phi(\theta)\right) \xrightarrow{d} \mathcal{N}(0, \phi'(\theta)\Sigma\phi'(\theta)^\top).$$

In particular,

$$\sqrt{n}\left(\phi(\hat{\theta}_n) - \phi(\theta)\right) \xrightarrow{d} \mathcal{N}(0, \phi'(\theta)\Sigma\phi'(\theta)^\top).$$

### Example: Sample Variance

Let  $X_i \in \mathbb{R}$ , i.i.d., with finite 4th moment. Let  $\hat{\sigma}^2 = \frac{1}{n} \sum (X_i - \bar{X})^2$ . What is  $\sqrt{n}(\hat{\sigma}^2 - \sigma^2) \xrightarrow{d} ?$

**Ans:** Note  $\hat{\sigma}^2 = \phi(\bar{X}, \bar{X}^2)$ , where  $\phi(x, y) = y - x^2$ . By CLT,

$$\sqrt{n} \left( \begin{pmatrix} \bar{X} \\ \bar{X}^2 \end{pmatrix} - \begin{pmatrix} \alpha_1 \\ \alpha_2 \end{pmatrix} \right) \xrightarrow{d} \mathcal{N} \left( \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \alpha_2 & \alpha_3 - \alpha_1\alpha_2 \\ \alpha_3 - \alpha_1\alpha_2 & \alpha_4 - \alpha_2 \end{pmatrix} \right)$$

where  $\alpha_k = \mathbb{E}[X^k]$ . Using  $\phi'(\theta) = (-2\alpha_1, 1)$  and applying the Delta Method yield

$$\sqrt{n}(\hat{\sigma}^2 - \sigma^2) \xrightarrow{d} \mathcal{N} \left( 0, \phi'(\theta) \Sigma \phi'(\theta)^\top \right).$$

What about  $\frac{1}{n-1} \sum (X_i - \bar{X})^2$ ? Simply rewrite

$$\frac{1}{n-1} \sum (X_i - \bar{X})^2 = \frac{n}{n-1} \hat{\sigma}^2 \approx \hat{\sigma}^2,$$

which asymptotically stays the same.

### Procedure

1. Write statistics as a function of simple statistics where we can apply the CLT.
2. Apply the Delta Method.

### Confidence Intervals for $\sigma$ :

Let

$$\hat{\alpha}_k = \frac{1}{n} \sum X_i^k, \quad k = 1, \dots, 4.$$

Then

$$\hat{\gamma}^2 = (-2\hat{\alpha}_1, 1) \hat{\Sigma} (-2\hat{\alpha}_1, 1)^\top,$$

and

$$\hat{\sigma}^2 \pm 1.96\hat{\gamma}/\sqrt{n}$$

is the asymptotic 95% confidence interval.

### Example:

$$T_n = \left( \hat{\sigma}^2, \frac{\bar{x}}{\hat{\sigma}} \right)$$
$$\sqrt{n} \left( T_n - \left( \sigma^2, \frac{\mu}{\sigma} \right) \right) \xrightarrow{d} ?$$